

1 **A fully remote randomized controlled trial of an ultra-brief digital**
2 **meditation intervention reduces internalizing symptoms**

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20 Abstract

21 Background: Scalable, low-burden behavioral interventions are needed to address rising subclinical
22 mental health symptoms. However, few randomized controlled trials have evaluated ultra-brief, remotely
23 delivered, meditation using multimodal outcome assessment under real-world conditions.

24
25 Methods: We conducted a fully remote randomized controlled trial (ClinicalTrials.gov: NCT06014281)
26 evaluating a focused-attention meditation intervention delivered via brief instructor training and
27 independent daily practice. A total of 299 meditation-naïve adults were randomized to immediate
28 intervention or waitlist control in a delayed-intervention design. Participants practiced ≥ 10 minutes daily
29 for 8 weeks within a 16-week study. Outcomes included validated self-report measures, web-based
30 cognitive tasks, and wearable-derived physiological metrics.

31
32 Results: Across randomized and within-participant replication phases, the intervention was associated
33 with significant reductions in anxiety and mind wandering, with effects remaining stable during 8-week
34 follow-up. Improvements were greatest among participants with higher baseline symptom burden. Sleep
35 disturbance improved selectively among individuals with poorer baseline sleep. Secondary outcomes,
36 including rumination, perceived stress, social connectedness, and quality of life, also improved. Cognitive
37 performance showed modest improvements primarily among lower-performing participants. Resting
38 heart rate exhibited nominal reductions.

39
40 Conclusions: An ultra-brief, fully remote meditation intervention requiring 10 minutes per day was
41 associated with sustained improvements in psychological functioning and smaller, baseline-dependent
42 effects on cognition in a non-clinical population. These findings support digital delivery of low-dose
43 meditation as a scalable preventive mental health strategy.

46 Key Words

47 Meditation, Internet-administered Meditation Training, mHealth, Remote Training

48

49 Introduction

50 The COVID-19 pandemic has led to a sustained global increase in stress, anxiety, and depressive
51 symptoms across diverse populations, including individuals with no prior psychiatric history
52 (Bourmistrova et al., 2022; Hossain et al., 2020; Kumar & Nayar, 2021; Liu et al., 2020; Saltzman et al.,
53 2024; Shanbehzadeh et al., 2021). Large-scale epidemiological studies have documented persistent
54 elevations in psychological distress well beyond the acute phases of lockdowns, with many individuals
55 reporting subclinical yet functionally impairing levels of stress and anxiety (COVID-19 Mental Disorders
56 Collaborators, 2021; Pierce et al., 2020; Tournon et al., 2025). Although these symptoms do not meet
57 diagnostic criteria for psychiatric disorders, they are associated with reduced well-being, impaired sleep,
58 decreased productivity, and increased risk of future psychopathology (Cuijpers & Smit, 2004;
59 Konstantopoulou et al., 2020). Because individuals with subthreshold symptoms often fall outside
60 traditional treatment pathways, scalable preventive interventions are urgently needed.

61
62 Meditation and related attentional training practices represent promising low-cost approaches for
63 improving emotional regulation and reducing stress. Meta-analyses demonstrate moderate reductions in
64 anxiety, stress, and depressive symptoms across clinical and non-clinical populations (Goldberg et al.,
65 2022; Goyal et al., 2014; Khoury et al., 2013), potentially mediated by improvements in attentional
66 control and regulation of stress-related physiological processes (Fox et al., 2016; Sagar et al., 2012,
67 2015; Tang et al., 2015). However, most evidence derives from intensive in-person training programs or
68 time-demanding daily practices, limiting real-world accessibility and long-term adherence.

69
70 Digital mental-health interventions provide an opportunity to overcome these barriers. Mobile-health
71 approaches allow individuals to engage in behavioral interventions remotely, at their own pace, and with
72 minimal infrastructure (Firth et al., 2017; Torous et al., 2021). For stress and anxiety management in
73 particular, sustained daily practice may be more important than session intensity, suggesting that brief,
74 repeatable interventions could offer scalable benefits (Elbaz et al., 2021; Staiano et al., 2025). Yet few
75 randomized controlled trials have tested whether very short daily meditation practices, on the order of
76 minutes rather than tens of minutes, can produce measurable psychological or physiological effects.

77
78 The present study evaluates a focused-attention meditation practice commonly referred to as SOS
79 meditation. Historically SOS meditation is taught within a broader contemplative tradition but
80 implemented here as a secular attentional training exercise (Munjal et al., 2025; Singh, 2022). The
81 practice involves directing attention inward while maintaining relaxed awareness and repeating a neutral
82 word or phrase. Unlike many structured mindfulness programs, the technique requires minimal
83 instruction, no specialized posture, and no ongoing instructor involvement, making it suitable for remote
84 delivery and routine integration into daily life. From a theoretical perspective, the practice aligns with
85 neurocognitive models of mindfulness emphasizing inward attention, narrow attentional aperture, and
86 meta-awareness (Lutz et al., 2015). Repeated engagement with these attentional processes may, over
87 time, reduce rumination and improve awareness of ongoing cognitive states while imposing relatively
88 low cognitive effort, potentially supporting adherence among meditation-naïve individuals. Preliminary
89 reports suggest beneficial psychological effects in remote settings (Munjal et al., 2025), but its broader
90 cognitive and physiological impacts, particularly at brief daily durations, remain unclear.

91
92 To address methodological limitations of prior meditation research, including small samples,
93 uncontrolled designs, and reliance only on self-report outcomes (Goldberg et al., 2022; Van Dam et al.,
94 2018), we conducted a fully remote randomized controlled trial with delayed-intervention replication.
95 Participants were assigned to either an immediate intervention group or a waitlist control group that later
96 received the same training, enabling both between-group comparisons and within-participant replication.
97 We recruited 299 meditation-naïve adults to complete a 16-week study consisting of an 8-week

98 intervention and follow-up period. Participants practiced at least 10 minutes per day using a remotely
99 delivered training protocol.

100
101 To capture potential effects across multiple systems, we adopted a multimodal assessment framework.
102 Psychological outcomes included validated measures of generalized anxiety (GAD-7), sleep disturbance
103 (PSQI), and mind wandering (MWQ), along with secondary measures of perceived stress, depressive
104 symptoms, rumination, social connectedness, quality of life, and character strengths. Cognitive
105 performance was assessed using web-based executive function and working memory tasks (Stroop and 2-
106 back). Physiological indices included resting heart rate and heart rate variability derived from nightly
107 wearable recordings. This integrative design allowed us to examine whether brief daily meditation
108 practice produces convergent changes across subjective experience, cognitive performance, and
109 autonomic regulation.

110
111 We hypothesized that a brief, remotely delivered focused-attention meditation intervention would reduce
112 anxiety and stress-related symptoms and produce convergent improvements across psychological,
113 cognitive, and physiological domains. By testing a minimal-burden intervention under real-world
114 conditions, this study evaluates the feasibility and efficacy of ultra-brief daily meditation as a scalable
115 preventive mental-health strategy.

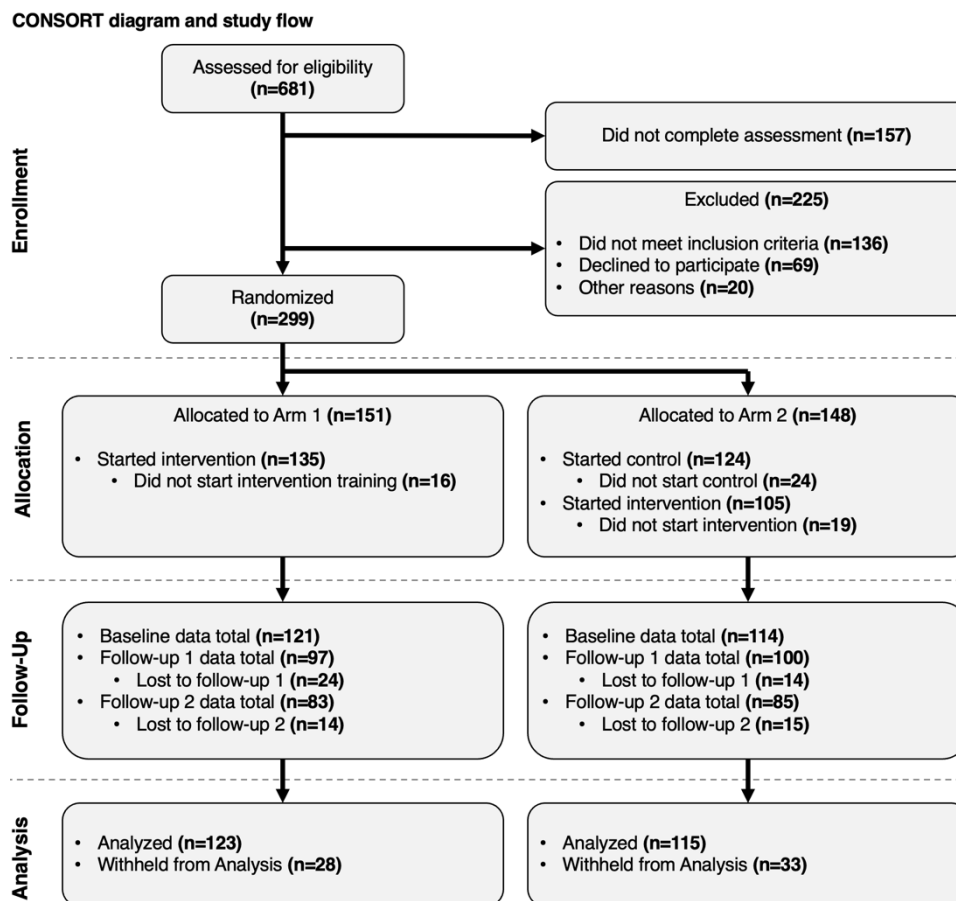
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117 Results

118 Study Population

119 A total of 681 individuals were assessed for eligibility, of whom 299 were randomized to either the
120 immediate intervention group (Arm 1, n=151) or the waitlist control group (Arm 2, n=148). Participant
121 flow, attrition, and analysis inclusion are summarized in the CONSORT diagram (Fig. 1).

122
123 The primary analysis sample had a mean age of 33.8 years and was 66.7% female. Baseline demographic
124 and clinical characteristics were similar across study arms (Table 1). Primary analyses adjusted for age
125 and sex.



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Fig. 1 | CONSORT flow diagram of participant recruitment, allocation, follow-up, and analysis. A total of 681 individuals were assessed for eligibility, of whom 299 were randomized to either the immediate intervention arm (n=151) or the waitlist control arm (n=148). The diagram summarizes exclusions, intervention initiation, follow-up completion, and inclusion in analysis for each arm.

Table 1 | Baseline characteristics of participants included in the primary analysis set. Values are mean (SD) or n (%). All analyses adjusted for age and sex.

<i>Characteristic</i>	<i>Arm 1 (n=123)</i>	<i>Arm 2 (n=115)</i>	<i>Combined (n=238)</i>
<i>Age (Mean, SD)</i>	33.03 (11.62)	34.69 (14.23)	33.83 (12.95)
<i>Female, n (%)</i>	71 (60.2%)	83 (73.5%)	154 (66.7%)
<i>Baseline GAD-7 (Mean, SD)</i>	6.46 (5.19)	5.50 (4.35)	6.00 (4.82)
<i>Baseline PSQI (Mean, SD)</i>	8.20 (2.38)	7.76 (2.35)	7.99 (2.37)
<i>Baseline MWQ (Mean, SD)</i>	17.39 (5.41)	16.57 (5.41)	16.99 (5.41)

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138

139 Anxiety and mind wandering decrease following brief daily meditation, particularly among those
140 with elevated baseline symptoms

141 *Anxiety decreased following meditation intervention and effects were sustained*

142 Daily meditation was associated with significant reductions in anxiety symptoms, as measured by the
143 GAD-7 (Fig. 2ai). During the randomized phase (T1–T2), participants assigned to the immediate
144 intervention arm exhibited greater reductions in anxiety relative to the waitlist control arm ($F=4.989$,
145 $p=0.038$; Table S1). When the waitlist group subsequently received the intervention (T2–T3), a
146 comparable within-participant reduction was observed ($t=-3.849$, $p<0.001$; Table S2), providing internal
147 replication of the effect. Anxiety reductions were maintained at follow-up, with no evidence of rebound
148 (Table S3). FDR correction was used in all analyses to control for multiple comparisons.

149
150 Pooling intervention periods across both arms revealed that participants with moderate or severe baseline
151 anxiety showed the largest improvements. Individuals with baseline GAD-7 scores (10–14) or (15–21)
152 demonstrated marked reductions (both $p<0.0001$; Table S4), with many shifting to lower clinical severity
153 categories following the intervention (Fig. 2a_{ii}). Participants with minimal baseline anxiety showed little
154 change.

155 *Sleep improved selectively among participants with baseline disturbance*

156 Across the three study points, changes in sleep disturbance (measured using the PSQI) did not reach
157 statistical significance (Fig. 2bi; Tables S1–S3). However, stratified analyses revealed significant
158 improvements among participants with third and fourth quartile baseline sleep disturbance ($p<0.05$; Fig.
159 2b_{ii}; Table S4), whereas individuals with first quartile baseline sleep showed somewhat higher PSQI
160 scores post-intervention ($p=0.042$; Table S4).

161
162 These findings suggest that sleep-related effects of brief meditation may be contingent upon baseline
163 symptom burden.

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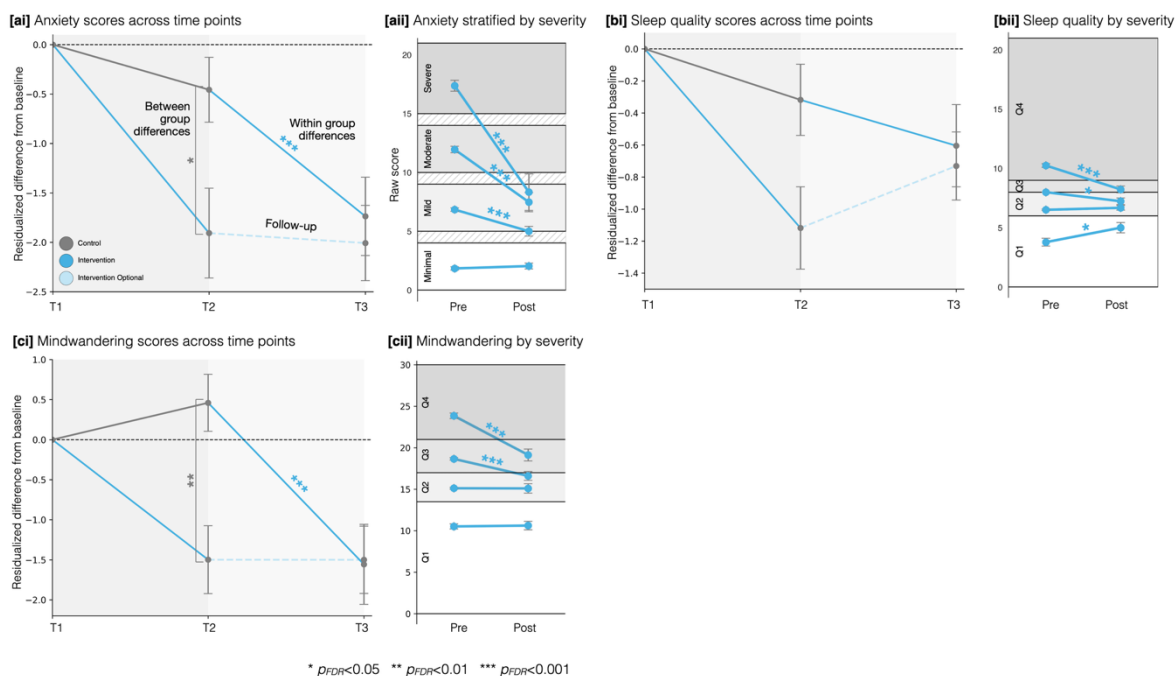
165 *Mind wandering decreased and effects were sustained*

166 Similar to anxiety scores, mind wandering (assessed using the MWQ) also declined following daily
167 meditation (Fig. 2ci). During the randomized phase, intervention participants showed greater reductions
168 relative to controls ($F=9.817$, $p=0.005$; Table S1). Replication was observed when the waitlist group
169 received training ($t=-4.146$, $p<0.001$; Table S2). Reductions were sustained eight weeks after the
170 intervention period.

171
172 Similar to anxiety, participants in the highest baseline quartiles of mind wandering exhibited the largest
173 decreases (Q3 and Q4, both $p<0.001$; Fig. 2c_{ii}; Table S4), whereas individuals with low baseline levels
174 showed no change.

175
176 Across the three primary self-report psychological outcomes, meditation was associated with robust
177 reductions in anxiety and mind wandering, with sustained effects at follow-up. Sleep improvements were
178 baseline dependent. In all domains, individuals with higher baseline symptom severity exhibited the
179 largest gains.

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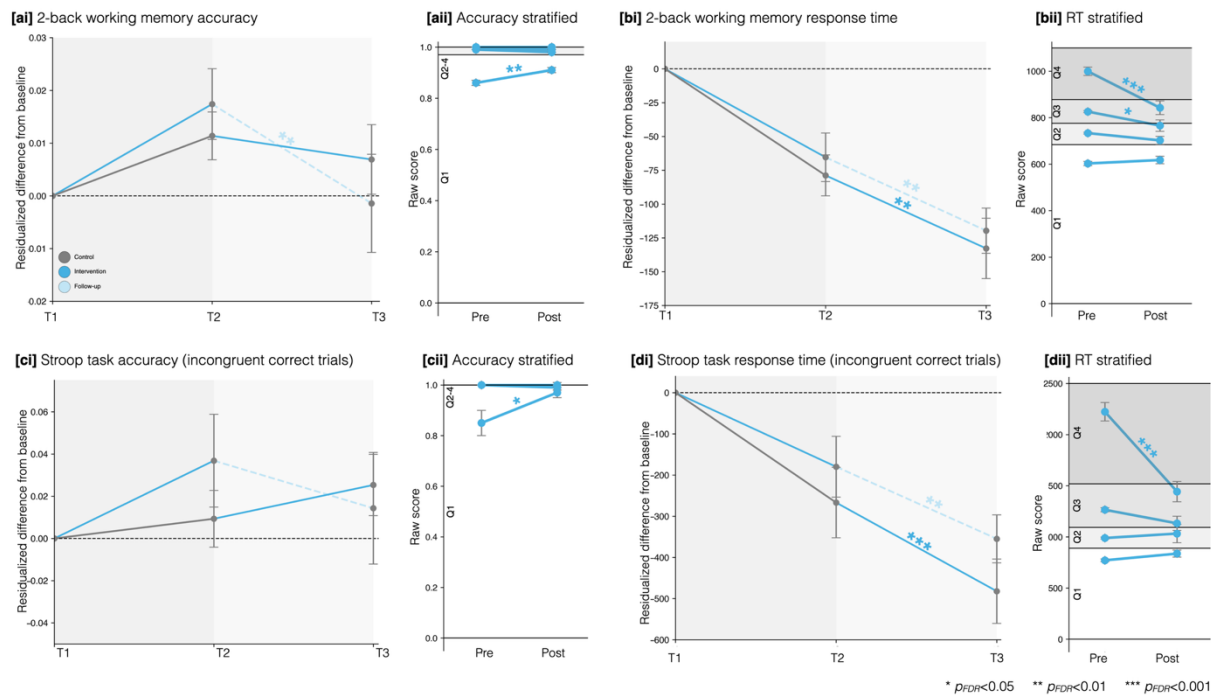
181
182 **Fig. 2 | Primary psychological outcomes.** [ai-ci] Residualized change from baseline in anxiety (GAD-7), sleep disturbance
183 (PSQI), and mind wandering (MWQ) across study timepoints. During the randomized phase (T1–T2), participants assigned to
184 the intervention arm showed greater reductions in anxiety and mind wandering relative to the waitlist control arm. Effects were
185 replicated when the waitlist group subsequently received the intervention (T2–T3) and were sustained at follow-up. Error bars
186 represent SEM. [aii-cii] Pre–post changes stratified by baseline severity are also presented. Participants with higher baseline
187 symptom levels exhibited larger reductions following the intervention, whereas individuals with minimal baseline symptoms
188 showed little change. Significance markers reflect FDR-adjusted tests.

189
190 **Cognitive improvements emerge primarily among lower-performing participants**

191
192 Cognitive performance was assessed using repeated administrations of the Stroop and 2-back working
193 memory tasks. At the whole-sample level, no significant between-arm differences in accuracy were
194 observed during the randomized phase (Fig. 3ai-di; Tables S9–S11). Response times decreased over time
195 across both intervention and control arms, consistent with task familiarity or practice effects, and did not
196 yield robust intervention-specific effects.

197
198 Baseline inspection revealed pronounced ceiling effects in accuracy across both tasks. To evaluate
199 potential benefits among individuals with greater room for improvement, we conducted stratified
200 analyses. Within the lowest-performing quartile, at baseline, participants exposed to meditation
201 demonstrated significant improvement in accuracy and reaction time on both the Stroop and 2-back tasks
202 (both $p < 0.05$; Fig. 3aii-dii; Table S12). This finding suggests that brief daily meditation may confer
203 modest cognitive benefits in participants with lower baseline executive performance, although robust
204 group-level effects were not observed.

205



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207
208 **Fig. 3 | Primary cognitive outcomes.** (ai-di) Residualized change from baseline in 2-back accuracy, 2-back reaction time,
209 Stroop incongruent accuracy, and Stroop incongruent response time across study timepoints. No robust between-arm differences
210 were observed at the whole-sample level, with improvements in reaction time occurring across both intervention and control
211 periods. Pre–post changes stratified by baseline performance quartile are also presented. (aii-dii) Participants in the lowest-
212 performing quartile at baseline exhibited significant improvements in accuracy and reaction time following meditation exposure,
213 whereas higher-performing participants showed minimal change. Error bars represent SEM. Significance markers reflect FDR-
214 adjusted tests.

215 Autonomic measures show directional but non-significant changes

216
217 Physiological indices were assessed using wearable-derived resting heart rate (RHR) and heart rate
218 variability (HRV) collected throughout the study period. During the randomized phase, RHR showed a
219 consistent downward trend among participants engaged in meditation relative to the waitlist control arm;
220 however, between-group differences did not reach statistical significance (Fig. S1; Tables S13–S16).
221 When the waitlist group subsequently received the intervention, an insignificant directional reduction in
222 RHR was observed. Among initial intervention participants, reductions were maintained through follow-
223 up.

224
225 Similarly, no significant group-level changes were detected for HRV across randomized or within-group
226 phases. HRV exhibited substantial inter-individual variability and did not differentiate intervention from
227 control periods.

229 Broad improvements across connectedness, rumination, stress, and quality of life

230
231 Secondary psychological measures showed patterns broadly consistent with primary outcomes (Fig. 4).
232 Although between-arm differences at week 8 did not survive FDR correction, nominal trends
233 (uncorrected $p < 0.05$) were observed for social connectedness, rumination, and quality of life scores
234 (Table S5). Further, within-participant quartile-based analyses following meditation exposure also
235 revealed significant improvements across several domains (Tables S6–S8).
236

237 *Social connectedness*

238 Social connectedness increased during meditation periods and replicated in the delayed-intervention arm
239 ($t=3.729$, $p<0.001$; Table S6). Participants with lower baseline connectedness exhibited the largest gains
240 (Fig. 4ai-ii; Table S8).
241

242 *Rumination*

243 Rumination decreased following meditation exposure in both arms. Reductions were most pronounced
244 among individuals with elevated baseline rumination (Fig. 4bi-ii; upper quartiles; $p<0.001$; Table S8),
245 consistent with a shift in maladaptive repetitive thought patterns.
246

247 *Perceived stress*

248 Perceived stress showed directional reductions across arms, reaching significance during the delayed-
249 intervention phase (Fig. 4ci-ii; $t=-2.496$, $p=0.023$; Table S6), suggesting modest stress-related benefits.
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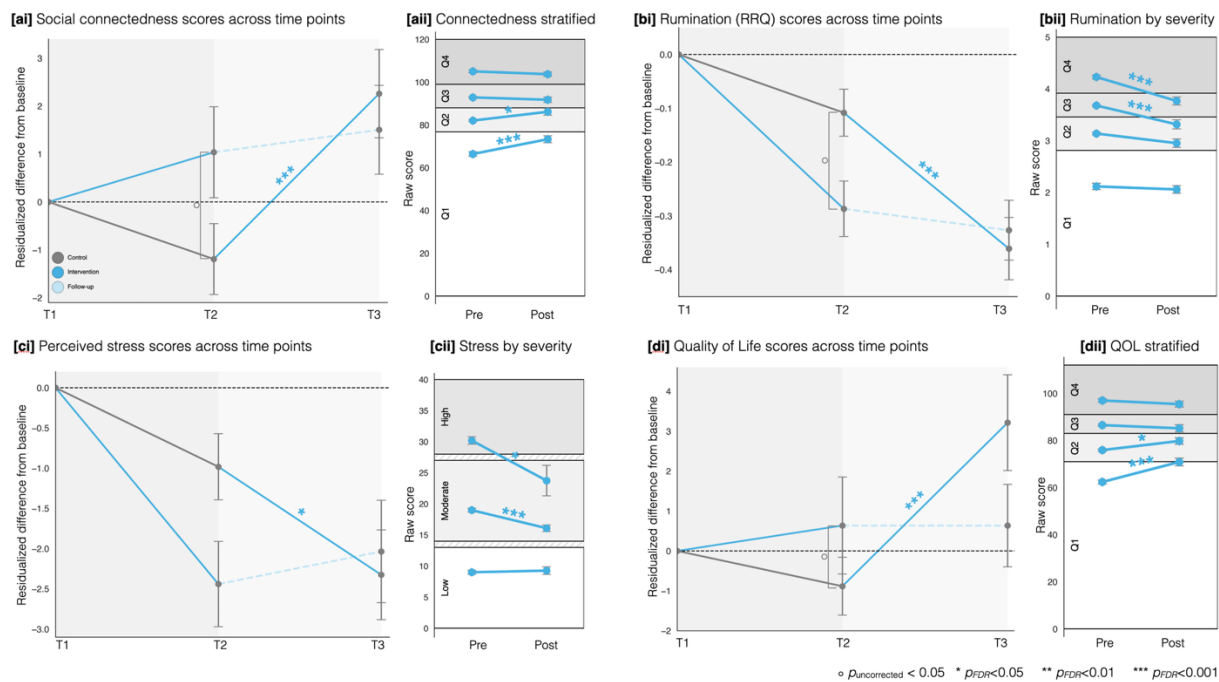
251 *Quality of life*

252 Quality of life improved following meditation, particularly among participants reporting lower baseline
253 well-being (Fig. 4di-ii; worst quartile; $t=5.335$, $p<0.001$; Table S8).
254

255 The remaining secondary psychological measures showed a more selective and baseline-dependent
256 pattern. Depressive symptoms improved mainly among participants with elevated baseline PHQ-8 scores,
257 and the worst quartile for reflection showed an increase following meditation exposure, whereas the BSS-
258 12 subscales exhibited mixed changes that were not consistently supported at the quartile level after
259 correction for multiple comparisons.

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261 Complete statistical details are provided in Supplementary Tables S5–S8.

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Fig. 4 | Secondary psychological outcomes. (a-d) Residualized change from baseline in secondary self-report measures, including social connectedness, rumination, perceived stress, and quality of life, across study timepoints. Group-level effects were directionally consistent with primary outcomes but generally smaller in magnitude and nominally significant group differences (uncorrected $p < 0.05$). Pre-post changes stratified by baseline severity or quartiles are also presented. Participants with lower baseline connectedness, higher rumination or stress, and lower quality of life exhibited the largest improvements following meditation exposure, whereas participants with minimal baseline impairment showed little change. Error bars represent SEM. Significance markers indicate FDR-adjusted tests.

273 Summary of univariate analysis

274 Across the psychological, cognitive, and physiological domains, the most robust and sustained effects
275 were observed in primary psychological outcomes, with baseline-dependent effects in cognition and
276 modest, non-significant trends in physiological measures. A summary of the primary imputed univariate
277 results is provided in Table 2; corresponding non-imputed results are reported in Table S17. The overall
278 pattern of findings was consistent with or without imputation.
279

280 **Table 2. Consolidated summary of univariate analysis.** Intervention effects across all the primary and secondary outcomes.
 281 Between-arm (T2; Week 8) reflects the randomized comparison between intervention and waitlist control groups. Replication
 282 reflects within-arm change in the delayed-intervention group following meditation exposure. Follow-up reflects stability of
 283 change from Week 8 (T2) to Week 16 (T3) in the initial intervention arm. Greater effect in higher-severity participants reflects
 284 baseline-stratified analyses among participants in the highest symptom quartile or clinically elevated range. Significance markers
 285 denote FDR-corrected p-values.

	<i>Outcomes</i>	<i>Between Arm (Week 8)</i>	<i>Replication</i>	<i>Follow-up</i>	<i>Greater effect in higher-severity participants</i>
<i>Primary Psychological Outcomes</i>	Anxiety (GAD-7)	*	***	☑ Sustained	***
	Mind Wandering (MWQ)	**	***	☑ Sustained	***
	Sleep (PSQI)	--	--	--	***
<i>Primary Cognitive Outcomes</i>	2-Back Accuracy	--	--	--	**
	2-Back Reaction Time	--	**	--	***
	Stroop Accuracy	--	--	--	*
	Stroop Reaction Time	--	***	--	***
<i>Primary Physiological Outcomes</i>	Resting Heart Rate	--	--	--	--
	Heart Rate Variability	--	--	--	--
<i>Secondary Psychological Outcomes</i>	Social Connectedness (SCS)	○	***	☑ Sustained	***
	Rumination (RRQ)	○	***	☑ Sustained	***
	Stress (PSS)	--	*	--	*
	Quality of Life (QOL)	○	***	☑ Sustained	***

○ $p_{\text{uncorrected}} < 0.05$ * $p_{\text{FDR}} < 0.05$ ** $p_{\text{FDR}} < 0.01$ *** $p_{\text{FDR}} < 0.001$

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289 Multivariate Structure of Self-Reported Measures

290 *Self-reported measures are organized along two latent dimensions*

291 To characterize shared variance across self-reported psychological measures, we performed exploratory
 292 factor analysis (EFA) on baseline questionnaire scores. A two-factor solution emerged, accounting for a
 293 substantial proportion of total variance (42%) (Fig 5a; Fig. S2; Table S18). Factor loadings were
 294 thresholded at 0.40 to enhance interpretability, consistent with established recommendations for
 295 interpretability (Brooks & Stevens, 1994), and factors were named based on the highest-loading
 296 measures.

297 The first factor, termed *Internalization and Negative Affect*, was defined by strong positive loadings from
 298 perceived stress, depressive symptoms, anxiety, rumination, and mind wandering, and negative loadings
 299 from quality of life and social connectedness. The second factor, termed *Resilience and Strength*, was
 300 characterized by positive loadings from intellectual, interpersonal, and temperance character strengths, as
 301 well as reflection. The two factors were moderately negatively correlated, indicating related but distinct
 302 psychological dimensions ($r = -0.54, p < 0.05$; Fig. 5b).

303 *Meditation selectively reduces the latent internalization dimension*

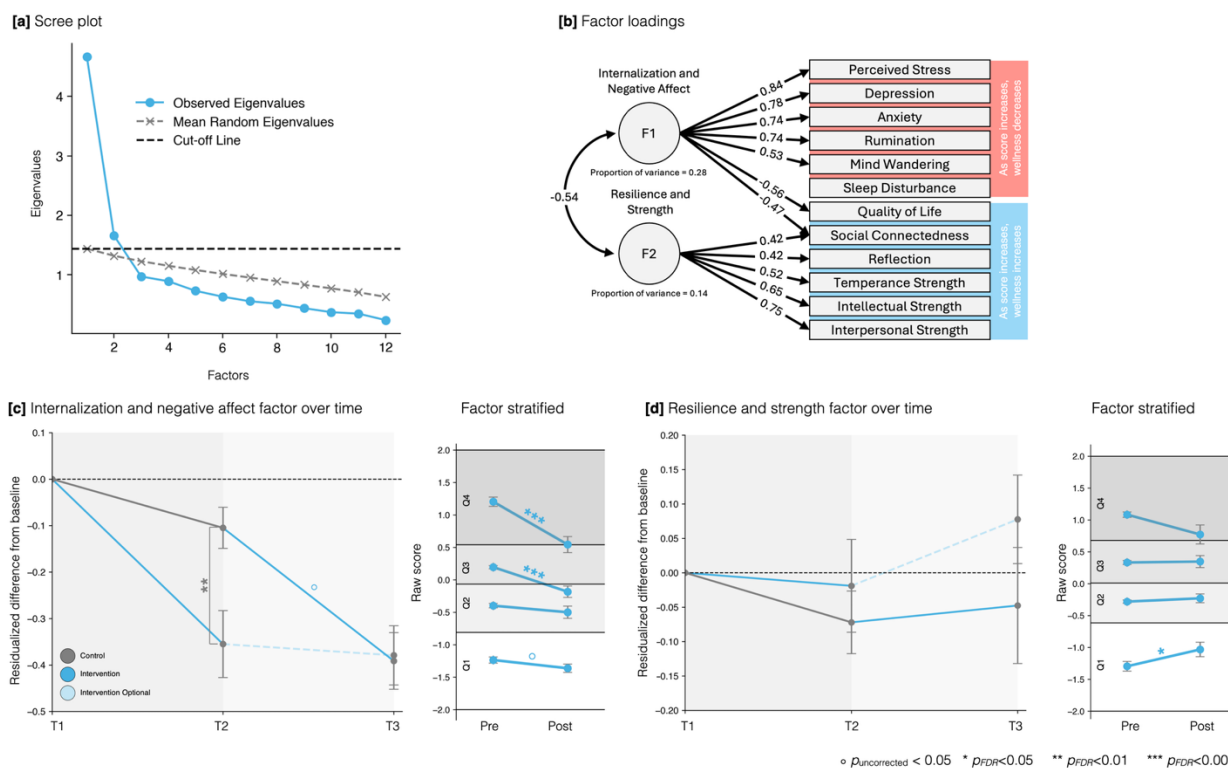
304 Factor scores derived from the baseline-fitted EFA model were examined across study timepoints using
 305 the same analytic framework as univariate outcomes. For the Internalization and Negative Affect factor,
 306 we observed a significant reduction during the randomized phase (T1–T2) in the intervention arm relative
 307 to control ($F = 11.195, p = 0.002$; Table S19), with consistent within-participant replication in Arm 2 during
 308 the delayed intervention period (T2–T3; $t = -5.106, p < 0.001$; Table S20) and sustained effects in Arm 1 at

309 follow-up. These results indicate a robust shift along the latent symptom axis of internalization following
 310 meditation (Fig. 5c). Baseline-stratified analyses revealed that participants with the highest symptom
 311 burden benefited the most from meditation intervention (Table S22).

312 In contrast, changes in the Resilience and Strength factor were not significant at the whole-sample level
 313 and did not show consistent between-arm differences during the randomized phase. However, baseline-
 314 stratified analyses revealed differential effects, with participants in the lowest baseline quartile exhibiting
 315 significant increases in resilience-related scores following intervention, whereas higher-baseline groups
 316 showed no change ($t=2.281, p=0.02$; Table S22).

317 Together, these findings suggest that the multivariate effects of brief daily meditation are primarily
 318 expressed as reductions in internalizing symptomatology, with more selective and baseline-dependent
 319 changes in resilience-related traits.

320



321
 322 **Fig. 5 | Exploratory factor analysis and latent change following meditation.** (a) Scree plot and parallel analysis of baseline
 323 (T1) self-report measures indicating a two-factor solution. The dashed line represents mean eigenvalues derived from 1,000
 324 random permutations. (b) Factor loadings ($|loading| \geq 0.40$) for the two-factor model fitted on baseline data. Factor 1
 325 (Internalization and Negative Affect) and Factor 2 (Resilience and Strength) were moderately negatively correlated ($r = -0.54$).
 326 (c) Change in Internalization and Negative Affect factor scores across study timepoints. Residualized change from baseline is
 327 shown for Arm 1 (intervention) and Arm 2 (control), with within-participant replication (Arm 2, T2–T3) and follow-up (Arm 1,
 328 T2–T3) phases. Right panel shows pre–post changes stratified by baseline factor quartiles, demonstrating greater reductions
 329 among participants with higher baseline scores. (d) Change in Resilience and Strength factor scores across study timepoints,
 330 displayed analogously to (c).

331

332 Baseline latent factors associated with individual response across domains

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334 Although meditation reduced Internalization and Negative Affect at the group level, substantial inter-
 335 individual variability in change scores was observed. To examine predictors of treatment response, we

336 computed partial correlations between baseline factor scores and post-pre change in individual outcomes,
337 controlling for age and sex. Analyses were conducted separately across psychological, physiological, and
338 cognitive domains with FDR correction applied within each domain.

339 *Associations with psychological outcomes*

340 Higher baseline Internalization and Negative Affect scores were associated with larger reductions in
341 anxiety ($r=-0.245, p=0.0123$) and larger increases in quality of life ($r=0.291, p=0.002$) and social
342 connectedness ($r=0.310, p=0.002$; Table S23). Associations with perceived stress and depressive
343 symptoms maintained a similar direction and were nominal at the raw level but did not survive FDR
344 correction.

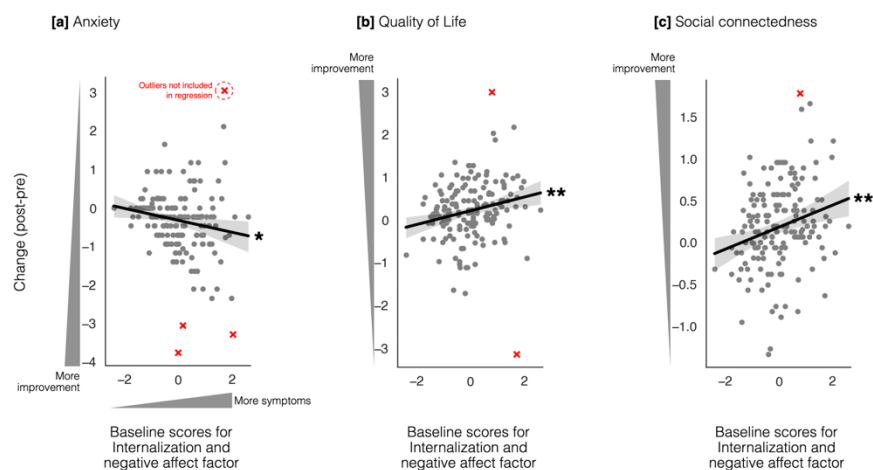
345
346 These findings indicate that individuals with elevated internalizing symptom burden at baseline
347 experienced greater psychological benefit from brief daily meditation practice.

348 Baseline Resilience and Strength scores showed a different pattern. Higher resilience was associated with
349 reductions in social connectedness following the intervention ($r = -0.250, p = 0.0121$).

350 *Associations with cognitive and physiological outcomes*

351 No significant associations were observed between baseline latent factor scores and changes in changes in
352 cognitive performance or physiological measures (Table S22-23).

353



354
355 **Fig. 6 | Baseline latent factors are associated with differential intervention response.** (a) Partial correlation coefficients
356 (controlling for age and sex) between baseline factor scores and post-pre (Δ) change in individual outcomes. Only associations
357 surviving FDR correction within the domain are shown. (b-c) Representative scatterplots illustrating significant associations
358 between baseline factor scores and Δ change in selected outcomes. Lines represent linear regression fits with 95% confidence
359 intervals. Stars indicate FDR-corrected significance levels as in other figures.

360 Discussion

361 This randomized, fully remote trial demonstrates that a brief (10 minutes/day) digitally delivered focused-
362 attention meditation intervention can produce measurable improvements in psychological functioning,
363 with effects concentrated along a latent internalizing symptom dimension and strongest among
364 individuals with elevated baseline symptom burden. By integrating validated self-report psychological
365 measures, web-based cognitive testing, and wearable-derived physiological indices within a scalable
366 remote design, this study advances prior internet-based mindfulness trials that have typically relied on
367 single-domain outcomes or lacked multimodal assessment (Devillers-Réolon et al., 2022; Sevilla-
368 Llewellyn-Jones et al., 2018; Simonsson & Marks, 2020).

369 Psychological Outcomes

370 Self-report measures demonstrated robust and consistent psychological benefits of brief regular
371 meditation practice. Participants in both study arms exhibited reductions in anxiety and mind wandering,
372 with the largest improvements observed among individuals who entered the study with moderate to
373 severe baseline symptom levels. Sleep disturbance showed a more nuanced pattern, with improvements
374 primarily evident among participants with the worst quartile of sleep difficulties, while individuals with
375 the best quartile of baseline sleep showed higher PSQI scores post-intervention. These findings are
376 consistent with prior work demonstrating that meditation interventions tend to be most effective for
377 individuals with elevated symptom burden (Kabat-Zinn et al., 1992; Krisanaprakornkit et al., 2006;
378 Orme-Johnson & Barnes, 2014).

379 Importantly, reductions in anxiety and mind wandering remained stable during the follow-up period, with
380 no evidence of symptom rebound. These sustained/stable effects suggest that brief SOS meditation may
381 induce changes in attentional and affective processes that generalize beyond periods of active practice.
382 The observed reductions in mind wandering align with prior neurocognitive and phenomenological work
383 linking meditation to enhanced meta-awareness and attentional control (Hasenkamp et al., 2012; Rahl et
384 al., 2017). For anxiety, the magnitude of change on the GAD-7 fell within ranges identified as minimally
385 clinically important, underscoring the clinical relevance of the findings beyond statistical significance.

386 These univariate improvements were mirrored at the latent level, where meditation was associated with a
387 reduction along a shared Internalization and Negative Affect dimension.

388 Cognitive Outcomes

389 Cognitive effects of a brief daily meditation were more modest. At the whole-sample level, no robust
390 between-group differences were observed on Stroop or 2-back task performance, likely reflecting ceiling
391 effects in this high-functioning, non-clinical sample. However, among participants in the worst-
392 performing quartile at baseline, meditation was associated with improvements in accuracy and processing
393 speed, particularly on Stroop incongruent trials. These findings are consistent with recent meta-analytic
394 work suggesting small but detectable benefits of mindfulness-related practices on working memory and
395 cognitive control, especially among individuals with lower baseline performance (Zainal & Newman,
396 2024).

397 The modest effect sizes observed here are also consistent with the broader literature, which remains
398 mixed regarding meditation's impact on cognition (Demnitz-King et al., 2023; Zainal & Newman, 2024).
399 Several factors may have constrained detectable cognitive effects in the present study, including the brief
400 duration of the intervention, individual variability in cognitive strategies, and limitations inherent to
401 remote testing environments (e.g., device heterogeneity, learning effects from repeated task exposure).
402 Nonetheless, the subgroup-specific improvements observed here suggest that baseline performance may
403 constrain detectable group-level effects and influence responsiveness to brief attentional training.

404 Physiological Outcomes

405 Physiological measures derived from wearable devices revealed directionally consistent changes. Resting
406 heart rate showed small reductions during meditation periods, with trends persisting into follow-up,
407 though these effects did not reach statistical significance at the whole-sample level. While non-
408 transcendental meditation practices have been associated with cardiovascular benefits in prior work (Shi
409 et al., 2017), the predominantly healthy nature of the present sample and the brevity of the intervention
410 likely limited the magnitude of detectable autonomic effects. Studies in clinical or higher-risk populations
411 may reveal stronger physiological responses. Lastly, wearable-derived measures, while scalable, may
412 lack sensitivity to subtle autonomic shifts induced by ultra-brief interventions.

413 **Multivariate Perspective and Individual Differences**

414 Exploratory factor analysis revealed that participants' self-reported experiences could be parsimoniously
415 represented along two latent dimensions: Internalization and Negative Affect, and Resilience and
416 Strength. Meditation was associated with a significant reduction along the Internalization and Negative
417 Affect axis, reflecting improvements across anxiety, depression, stress, rumination, and mind wandering.
418 Notably, baseline factor scores associated with intervention responsiveness: individuals with higher
419 Internalization and Negative Affect showed greater psychological improvement. These findings suggest
420 that brief digital meditation does not exert uniform effects but instead shifts individuals along a
421 dimensional symptom continuum, with baseline position influencing magnitude and domain of response.

422 This pattern aligns with prior evidence suggesting that baseline characteristics meaningfully shape
423 response to mindfulness interventions, with greater gains often observed among those with lower baseline
424 well-being or mindfulness (Vergara et al., 2022). Together, these findings highlight the importance of
425 personalized approaches to meditation-based interventions, where baseline psychological profiles may
426 help identify individuals most likely to benefit across different outcome domains.

427 Because baseline symptom levels were correlated with magnitude of change, regression-to-the-mean
428 effects cannot be fully excluded. However, the presence of randomized between-arm effects during the
429 initial intervention phase and replication of improvements in the delayed intervention arm provide
430 support for intervention-related change beyond statistical artifact.

431 **Limitations and Future Directions**

432 Several limitations should be considered when interpreting the findings of this study.

433 *Attrition and engagement.* Attrition was substantial, with many participants failing to complete all study
434 procedures. Commonly reported reasons included competing demands, loss of interest, and difficulty
435 maintaining engagement over the study duration. High dropout rates are a well-documented challenge in
436 internet-based meditation and mindfulness trials (Krisanaprakornkit et al., 2006) and underscore that self-
437 directed meditation may not be a universally suitable intervention, particularly for individuals with lower
438 intrinsic motivation or limited capacity for sustained self-regulation. As such, the present findings may
439 preferentially reflect outcomes among participants who were able to remain engaged with a self-guided
440 intervention.

441 *Challenges inherent to fully remote study designs.* While the fully remote nature of the trial enabled
442 scalability and ecological validity, it also introduced several sources of data loss and variability.
443 Participants frequently experienced difficulties with survey completion, wearable adherence, and
444 connectivity or synchronization issues with Fitbit devices, resulting in a reduced number of complete
445 physiological datasets. Similarly, completion rates for web-based cognitive tasks were lower than
446 anticipated, reflecting both participant burden and the learning curve associated with unfamiliar testing
447 platforms. In addition, remote cognitive testing is inherently sensitive to uncontrolled environmental
448 factors (e.g., distractions, device heterogeneity, internet latency), which may have increased measurement
449 noise and attenuated detectable effects.

450 *Ceiling effects and task sensitivity.* The cognitive tasks employed (Stroop and 2-back) exhibited
451 pronounced ceiling effects in this high-functioning, non-clinical sample, limiting sensitivity to
452 intervention-related changes at the whole-sample level. While subgroup analyses revealed improvements
453 among participants with lower baseline performance, future studies may benefit from adaptive or more
454 challenging paradigms to better capture cognitive change in non-clinical populations.

455 *Lack of an active control condition.* The use of a waitlist control limits causal specificity, as expectancy
456 effects, demand characteristics, or nonspecific engagement effects cannot be fully ruled out. Inclusion of
457 an active control condition (e.g., health education, relaxation exercises, or non-meditative attention tasks)
458 would strengthen causal inference and may also improve retention in control arms by providing

459 participants with a structured daily activity. The absence of such a control represents an important
460 limitation and should be addressed in future trials.

461 *Constraints on crossover and washout designs.* Because meditation is a self-dosed behavioral
462 intervention with no clear biological half-life, it was not feasible to implement a traditional crossover
463 design with a washout period. Consequently, residual or carryover effects cannot be fully disentangled
464 from sustained intervention effects, particularly for outcomes that persisted during follow-up.

465 *Sample size and statistical power.* Although the study enrolled nearly 300 participants, attrition and
466 incomplete datasets reduced the effective sample size for several analyses, particularly for physiological
467 and cognitive outcomes. As a result, the study may have been underpowered to detect small-to-moderate
468 effects in these domains, especially given the use of conservative multiple-comparisons corrections.

469 *Generalizability.* The sample consisted exclusively of undiagnosed adults, many of whom resided in or
470 near the San Francisco Bay Area. This limits generalizability to clinical populations, individuals with
471 limited access to technology, or populations with different cultural or socioeconomic backgrounds. Future
472 studies should incorporate broader recruitment strategies and explicitly test SOS meditation in clinically
473 diagnosed samples to assess translational relevance.

474 Conclusions

475 Despite these limitations, this study demonstrates that ultra-brief, remotely delivered focused-attention
476 meditation can meaningfully reduce internalizing symptoms in a non-clinical population. Effects were
477 most robust for anxiety and mind wandering and were concentrated among individuals with elevated
478 baseline symptom burden, consistent with a dimensional model of intervention responsiveness. While
479 cognitive and physiological effects were more modest, findings highlight the feasibility of integrating
480 low-dose meditation into scalable digital mental health frameworks. Together, these results support the
481 potential of brief, internet-delivered meditation as a preventive strategy within digitally mediated models
482 of care, particularly when personalized to baseline psychological profile.
483

484 Methods

485 Study Design

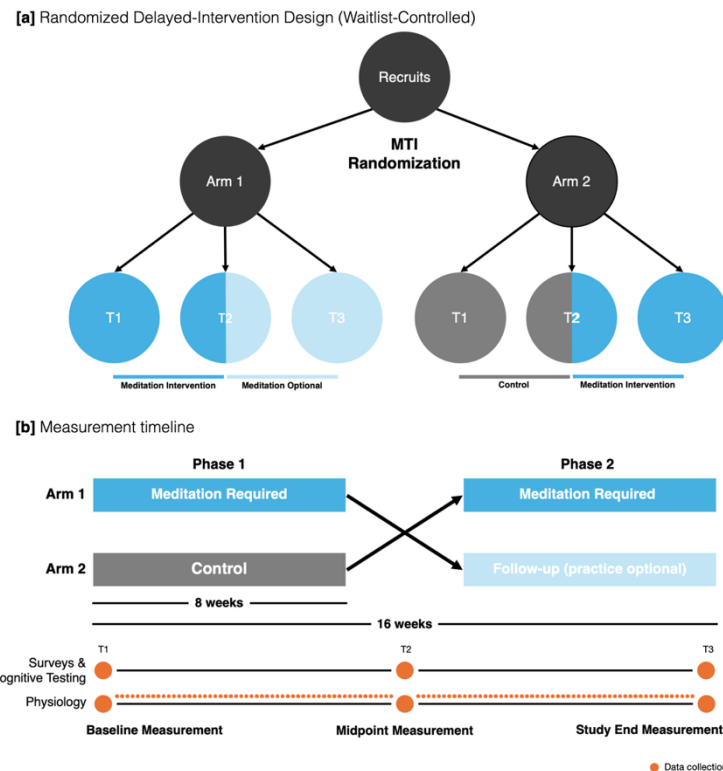
486 This study was a 16-week, fully remote randomized controlled trial with a delayed-intervention (waitlist-
487 controlled) design (Fig. 7). The active meditation intervention lasted 8 weeks, followed by an 8-week
488 follow-up period. The study was conducted between December 2023 and July 2025, with coordination
489 and oversight provided by staff at Stanford University School of Medicine (California, USA).
490

491 Participants assigned to Arm 1 initiated the meditation intervention immediately following baseline
492 assessment and completed 8 weeks of daily practice. After week 8, they entered a naturalistic follow-up
493 phase during which continued practice was optional. Participants assigned to Arm 2 completed an initial
494 8-week waitlist control period prior to receiving the identical 8-week meditation intervention. This design
495 enabled between-arm comparisons during the randomized phase (weeks 0–8) and within-participant
496 replication in the delayed-intervention arm (weeks 8–16).
497

498 Primary self-report and cognitive outcomes were assessed at baseline (week 0), post-intervention (week
499 8), and follow-up (week 16). Physiological measures were collected daily throughout the study.
500 Intervention adherence was monitored during each participant's active meditation phase.
501

502 Because meditation is a behavioral intervention without a defined pharmacokinetic half-life, no formal
503 washout period was implemented, and intervention and follow-up phases proceeded consecutively.
504

505 The trial was registered at ClinicalTrials.gov (Identifier: NCT06014281).
506



507
508 **Fig. 7 | Study design and assessment timeline.** (a) Participants were randomized 1:1 to either immediate intervention (Arm 1)
509 or waitlist control with delayed intervention (Arm 2). Data were collected at three timepoints (T1: baseline; T2: week 8; T3:
510 week 16). For Arm 1, T1–T2 corresponds to the intervention period and T2–T3 to follow-up. For Arm 2, T1–T2 corresponds to
511 the waitlist control period and T2–T3 to the intervention period. (b) Study timeline illustrating the 16-week structure. Self-report
512 and cognitive assessments were administered at T1, T2, and T3. Physiological measures were collected daily. No washout period
513 was implemented.
514

515 Participant Recruitment and Eligibility

516 Participants were recruited across the United States between September 2023 and December 2024 via
517 email distribution lists, social media postings, clinical trial registries, and word-of-mouth outreach.
518 Although participation was open nationally, most enrolled participants resided in the San Francisco Bay
519 Area.

520
521 Eligible participants were adults aged ≥ 18 years who expressed interest in learning meditation and were
522 willing to complete a fully remote 16-week study. Exclusion criteria included any current or recent
523 (within six months) psychiatric or neurological diagnosis; medical or sensory conditions that could
524 interfere with study participation or outcome measurement (e.g., significant visual impairment including
525 color blindness, hearing impairment, diagnosed sleep disorders such as narcolepsy, recent prolonged
526 hospitalization); inability to wear a wrist-based fitness tracker; and prior regular meditation practice,
527 defined as meditating ≥ 3 times per week for >10 minutes per session within the six months preceding
528 enrollment.
529

530 Participants were required to reside in the United States, have reliable internet access, and be able to
531 comply with daily data collection procedures during intervention periods. Screening was conducted
532 online using structured questionnaires, and eligibility was confirmed by trained study staff prior to
533 enrollment. Participants received no financial compensation beyond device provision.

534
535 All participants provided informed consent electronically prior to study participation. Consent included
536 agreement to random assignment, repeated assessments, daily data collection during intervention phases,
537 and secure storage of de-identified data for research purposes. The study protocol was approved by the
538 Stanford University Institutional Review Board (IRB #68784) and registered at ClinicalTrials.gov
539 (NCT06014281).
540

541 **Randomization**

542 Participants were randomized in a 1:1 ratio to one of two study arms using the NIH Clinical
543 Randomization Tool implementing the Maximum Tolerated Imbalance (MTI) procedure (Berger et al.,
544 2003). The maximal imbalance constraint was set to three participants to maintain near-equal allocation
545 throughout enrollment. Randomization was not stratified by age or sex.
546

547 Allocation sequences were generated prior to study initiation. Participants were assigned sequentially as
548 they completed eligibility screening and provided informed consent, allowing balanced allocation during
549 rolling enrollment. Study staff had no influence over allocation sequence
550

551 **Meditation Intervention**

552 The intervention consisted of three components: (1) an initial remote training session, (2) a mid-
553 intervention check-in session, and (3) daily self-guided meditation practice over an 8-week period.

554 *Initial Training Session*

555 Participants attended a 60-minute live remote training session conducted via secure videoconferencing.
556 Sessions were delivered in small groups (≤ 12 participants) by a single instructor (author M.S.) using a
557 standardized script to ensure consistency. When six or more participants were present, a second trained
558 staff member (authors C.G. or S.P.) joined to maintain a maximum 6:1 participant-to-staff ratio.
559

560 The training included an overview of focused-attention meditation principles, two guided practice periods
561 (10 mins each), and structured discussion segments following each practice period. Participants were
562 required to keep cameras on to verify engagement and were encouraged to ask questions during the
563 session.

564 *Mid-Intervention Check-In*

565 Approximately four weeks after initiating practice, participants attended a 30-minute remote check-in
566 session led by trained study staff. Sessions were conducted in small groups (≤ 5 participants) to maintain
567 individualized engagement. The check-in included guided practice (10 min), troubleshooting of common
568 challenges, and reminders regarding study procedures. If scheduling conflicts arose, sessions could be
569 rescheduled within the intervention window.

570 *Daily Meditation Practice*

571 Participants were instructed to engage in a focused-attention meditation practice for a minimum of 10
572 minutes per day during the 8-week intervention period (56 days). The practice was derived from a
573 meditation approach historically associated with the organization Science of Spirituality (SOS) but was
574 implemented in this study as a secular attentional training exercise.
575

576 To minimize expectancy effects and preserve methodological neutrality, the intervention was presented
577 as a focused-attention meditation protocol without emphasis on spiritual framing.
578

579 Participants were provided with standardized meditation instructions (Singh, 2022), delivered verbatim
580 during training and reproduced below to ensure intervention transparency and reproducibility:

- 581 1. *Close your eyes gently and relax, as you would when preparing to sleep.*
- 582 2. *Keep your attention alert and avoid straining your eyes, crossing them, or directing them*
583 *upward.*
- 584 3. *Direct your inner gaze approximately 8–10 inches into the field of darkness in front of you along*
585 *the horizontal plane.*
- 586 4. *Silently repeat any calming word.*
- 587 5. *Remain still and observe internal experience calmly, as though watching images arise on a*
588 *screen.*

590 Participants were permitted flexibility in accumulating their daily practice time (e.g., one 10-minute
591 session or multiple shorter sessions (≥ 5 minutes) for a total of ≥ 10 minutes every day. Continued
592 meditation beyond the required 8-week period was optional during follow-up phases.

593

594 Outcome Measures

595 Outcomes were assessed across three primary domains: psychological self-report, cognitive performance,
596 and physiological regulation.

597 Primary Outcomes

598 Primary psychological outcomes included generalized anxiety (GAD-7), sleep disturbance (Pittsburgh
599 Sleep Quality Index), and mind wandering (Mind Wandering Questionnaire). These instruments are
600 validated measures widely used in clinical and non-clinical populations (Spitzer et al., 2006; Buysse et
601 al., 1989; Mrazek et al., 2013).

602
603 Primary physiological outcomes included resting heart rate (RHR) and heart rate variability (HRV),
604 derived from nightly recordings using Fitbit Charge 6 wearable devices. Daily measurements were
605 aggregated within predefined study windows for analysis.

606

607 Primary cognitive outcomes included performance on web-based Stroop and 2-back tasks administered
608 via the Gorilla.sc platform. From these tasks, accuracy and reaction time metrics were derived to index
609 executive function and working memory.

610

611 Secondary Outcomes

612 Secondary psychological measures included perceived stress (PSS), depressive symptoms (PHQ-8),
613 quality of life, social connectedness, rumination and reflection, and character strengths. Secondary
614 physiological outcomes included wearable-derived sleep quality metrics.

615

616 A detailed summary of all primary and secondary measures, including assessment timing and operational
617 definitions, is provided in Tables 3 and 4.

618

619 All outcome measures, study design, etc. were preregistered before data collection (ClinicalTrials.gov:
620 NCT06014281).

621

622

623

624

625

626 **Table 3: Primary Measures.** A summary table listing primary measures of interest, their description, and timepoints at which
 627 the data were collected.
 628

<i>Measure Type</i>	<i>Outcome Measure</i>	<i>Description</i>	<i>Time Frame</i>
<i>Self-report scores</i>	Change in Generalized Anxiety Disorder Questionnaire Score	Anxiety as measured using Generalize Anxiety Disorder Questionnaire	3 times: T1 (wk 0), T2 (wk 8), T3 (wk 16)
	Change in Pittsburgh Sleep Quality Index Questionnaire Score	Sleep Quality measured using Pittsburgh sleep quality index (PSQI)	3 times: wk 0, wk 8, wk 16
	Change in Mind Wandering Questionnaire Score	Mind wandering measured using the Mind Wandering Questionnaire	3 times: wk 0, wk 8, wk 16
<i>Physiological</i>	Change in Resting Heart Rate	Obtained daily from Fitbit Charge 6 Fitness Tracker worn Nightly	112 times: daily from wk 0 to wk 16
	Change in Heart Rate Variability	Obtained daily from Fitbit Charge 6 Fitness Tracker worn Nightly	112 times: daily from wk 0 to wk 16
<i>Cognitive Test</i>	Change in Stroop Test Score	Using the Stroop test, which measures the average reaction time between incongruent and congruent trials	3 times: wk 0, wk 8, wk 16
	Change in 2-back Test Score	Using the N-back test is a memory task where participants must remember letters from N trials ago.	3 times: wk 0, wk 8, wk 16

629
 630 **Table 4: Secondary Measures.** A summary table listing secondary measures of interest, their description, and timepoints at
 631 which the data were collected.
 632
 633

<i>Measure Type</i>	<i>Outcome Measure</i>	<i>Description</i>	<i>Time Frame</i>
<i>Self-report scores</i>	Change in Stress Scale Measure Questionnaire Score	Stress measured using Perceived Stress Scale measure	3 times: T1 (wk 0), T2 (wk 8), T3 (wk 16)
	Change in Depression Questionnaire Score	Depression measured using Patient Health Questionnaire	3 times: wk 0, wk 8, wk 16
	Change in Quality of Life Scale Measures Questionnaire Score	Quality of life measured using Quality of Life Scale measures	3 times: wk 0, wk 8, wk 16
	Change Social Connectedness Scale-Revised Questionnaire Score	Social Connectedness measured using Social Connectedness Scale-Revised	3 times: wk 0, wk 8, wk 16
	Change in Rumination Reflection Questionnaire Score	Rumination and reflection scales measured using Rumination Reflection Questionnaire	3 times: wk 0, wk 8, wk 16
	Change in Brief Strengths Scale Questionnaire Score	Strengths scale measured using Brief Strengths Scale measures individuals' Temperance Strength, Intellectual Strength, and Interpersonal Strength.	3 times: wk 0, wk 8, wk 16
<i>Physiological</i>	Change in Sleep Quality Measure	Obtained daily from Fitbit Charge 6 Fitness Tracker worn Nightly	112 times: daily from wk 0 to wk 16

634
 635
 636 **Data Collection**

637 All study data were collected remotely using secure, web-based platforms. Participants were provided
 638 with written instructions and contact information for study staff to address technical or procedural
 639 questions. Because several questionnaires assessed sensitive mental health domains, participants were
 640 permitted to skip individual items if desired.
 641

642 *Self-Report Measures*

643 Self-report data were collected using REDCap, a secure web-based data capture system. Demographic
 644 information was obtained at baseline to assess eligibility and characterize the sample. These data included
 645 age, sex, ethnicity, education level, employment status, marital status, and prior meditation experience
 646 and were not reassessed at follow-up.
 647

648 Validated mental health and well-being questionnaires were administered at baseline (T1), week 8 (T2),
649 and week 16 (T3). Measures included the Generalized Anxiety Disorder–7 (GAD-7), Social
650 Connectedness Scale–Revised (SCS-R), Pittsburgh Sleep Quality Index (PSQI), Rumination–Reflection
651 Questionnaire (RRQ), Quality of Life Scale (QOLS), Patient Health Questionnaire–8 (PHQ-8), Perceived
652 Stress Scale (PSS), Brief Strengths Scale–12 (BSS-12), and Mind Wandering Questionnaire (MWQ).

653
654 Responses were reviewed upon submission. Participants whose scores fell within clinically severe ranges
655 were contacted by trained study staff and provided with appropriate mental health resources according to
656 a predefined safety protocol.
657

658 *Meditation Adherence Monitoring*

659 During each participant’s 8-week intervention phase (Arm 1: T1–T2; Arm 2: T2–T3), daily meditation
660 logs were completed via REDCap. Participants recorded duration of practice, subjective focus, subjective
661 calm following practice, and date of completion. Participants reporting fewer than four days of meditation
662 within a given week were contacted to support adherence. Logs submitted outside the required 8-week
663 intervention window were not included in analyses.
664

665 *Physiological Data Collection*

666 Participants were provided with a Fitbit Charge 6 wearable device prior to study initiation. Devices were
667 worn nightly from approximately 10–15 minutes before sleep until 10–15 minutes after waking, yielding
668 up to 112 nights of data across the 16-week study. Although participants could wear the device during
669 daytime hours, only nocturnal data were analyzed.
670

671 Physiological data were aggregated using Fitabase, a third-party platform for wearable data extraction.
672 Resting heart rate (RHR) and heart rate variability (HRV) were extracted for analysis. Sleep-stage metrics
673 were collected but are not reported in the present manuscript.
674

675 *Cognitive Task Administration*

676 Cognitive performance was assessed at T1, T2, and T3 using web-based Stroop and 2-back tasks
677 delivered via the Gorilla.sc platform. Each task was administered in three runs per timepoint (2-back: 30
678 trials per run; Stroop: 27 trials per run). Across the full study, participants completed up to nine runs per
679 task.
680

681 To reduce variability associated with remote testing environments, occasional misunderstanding of task
682 instructions, or incomplete runs, the run with the highest accuracy at each timepoint was selected for
683 analysis. Task and run order were randomized across participants and sessions.
684

685 The Stroop task measured selective attention and interference control through accuracy and reaction time
686 differences between congruent and incongruent color-word trials. The 2-back task assessed working
687 memory by requiring participants to identify matches between the current stimulus and the letter
688 presented two trials earlier. Accuracy and reaction time were recorded for all trials.
689

690 *Data Analysis*

691 All analyses were conducted in accordance with the preregistered study design. Analyses were performed
692 using R (v4.2.0) and Python (v3.14.0). All tests were two-sided unless otherwise specified. False

693 discovery rate (FDR; Benjamini–Hochberg) correction was applied within predefined outcome families
694 where indicated.
695

696 *Handling of Missing Data*

697 Missing data were addressed using multiple imputation by chained equations (MICE; van Buuren &
698 Groothuis-Oudshoorn, 2011)) with predictive mean matching. Imputation was performed separately
699 within each outcome domain and analytic window to preserve the temporal structure of the study. Fifty
700 imputed datasets were generated using 20 iterations per dataset. Study arm and sex were not imputed; age
701 was imputed when missing.
702

703 For between-arm comparisons at week 8 (T1–T2), imputation models included baseline (T1) and post-
704 intervention (T2) scores, study arm, sex, and age. For within-arm analyses (replication in Arm 2 and
705 follow-up in Arm 1; T2–T3), imputation was restricted to the T2–T3 window, with week 8 and week 16
706 scores estimated, and baseline score, sex, and age included as auxiliary variables.
707

708 To preserve the randomized structure for Arm 2 at week 8, a post-imputation last-observation-carried-
709 forward (LOCF) rule was applied when week-8 values were missing but baseline values were observed.
710 Estimates were pooled across imputations using Rubin’s rules.
711

712 Analytic estimates (effects, standard errors, confidence intervals, and p-values) were pooled across
713 imputations using Rubin’s rules (Rubin, 1987), as implemented in the mice package.

714 *Primary Analytic Framework*

715 For all outcome domains, we implemented a three-part analytic strategy:

- 716 1. Discovery (between-group) analysis comparing Arm 1 (intervention) vs Arm 2 (control) during
717 the randomized phase (T1–T2)
- 718 2. Replication (within-group) analysis in Arm 2 during delayed intervention (T2–T3)
- 719 3. Sustainability analysis in Arm 1 during follow-up (T2–T3)

720 Age and sex were included as covariates in all adjusted between-group models.
721

722 *Self-Report Outcomes*

723 Between-group effects at week 8 were tested using ANCOVA:

$$725 \text{Outcome}_{T_2} \sim \text{Outcome}_{T_1} + \text{Arm} + \text{Age} + \text{Sex}$$

727 Within-arm changes (replication and follow-up) were assessed using paired-sample t-tests.
728

729 To evaluate differential effects by baseline severity, pre–post changes were pooled across intervention
730 periods (Arm 1: T1–T2; Arm 2: T2–T3). Participants were grouped by established clinical cutoffs (when
731 available) or quartiles. Paired t-tests were conducted within subgroups, with FDR correction applied
732 across subgroup comparisons.

733 *Physiological Outcomes*

734 Primary physiological outcomes included resting heart rate (RHR) and heart rate variability (HRV).
735 Between-group comparisons at week 8 were analyzed using ANCOVA models analogous to those used
736 for self-report outcomes. Within-arm replication and follow-up analyses were conducted using paired t-
737 tests.
738

739 Baseline-stratified analyses were performed using pooled intervention windows and quartile grouping,
740 with FDR correction applied within physiological outcomes.
741

742 *Cognitive Outcomes*

743 Cognitive outcomes were derived from Stroop and 2-back tasks. At each timepoint, the run with the
744 highest accuracy was selected for analysis to reduce variability associated with remote testing conditions.
745

746 Primary cognitive metrics included accuracy and reaction time for Stroop congruent and incongruent
747 trials and total accuracy and reaction time for 2-back trials. Analyses followed the same discovery,
748 replication, and follow-up framework described above, using ANCOVA for between-group comparisons
749 and paired t-tests for within-group analyses. Baseline-stratified analyses used quartile grouping due to the
750 absence of clinical cutoffs.
751

752 *Exploratory Factor Analysis*

753 Exploratory factor analysis (EFA) was conducted on baseline self-report measures. Factorability was
754 assessed using Bartlett's test of sphericity and the Kaiser–Meyer–Olkin statistic. Horn's parallel analysis
755 determined the number of factors retained (factor-analyzer 0.5.1; scikit-learn 1.1.1; scipy 1.13.1;
756 statsmodels 0.14.4).
757

758 A two-factor oblimin-rotated solution was estimated. Loadings $\geq |0.40|$ were retained for interpretability.
759 Factor scores were extracted from the baseline-fitted model (Stevens, 1992).
760

761 Between-group comparisons at week 8 were analyzed using ANCOVA models analogous to those used
762 for self-report outcomes. Within-arm replication and follow-up analyses were conducted using paired t-
763 tests.
764

765 *Associations Between Baseline Factors and Change*

766 To examine individual differences in responsiveness, partial correlations were computed between
767 baseline factor scores and standardized pre–post change scores within self-report, physiological, and
768 cognitive domains, controlling for age and sex. FDR correction was applied within each outcome domain.
769 Only associations surviving FDR correction were reported.
770

771 **Data availability**

772 De-identified data is available upon reasonable request from the corresponding author. Code is available
773 at the *Brain Dynamics Lab GitHub* page.
774

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778

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802 **M.S.** Conceptualization, Methodology, Investigation, Writing – Original Draft, Project Administration,
803 Supervision, Funding acquisition
804

805 Ethics declaration

806 This study was conducted according to the ethical guidelines outlined in the *Declaration of Helsinki*.
807 Approval was obtained, and the study was monitored by the Stanford Institutional Review Board. All
808 participants provided informed consent and were informed of the option to withdraw from the study at
809 any time without penalty. There are no financial interests to report.
810

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